

Vol. 40 (Number 7) Year 2019. Page 19

Modeling of the cycle of reproduction process in the agrarian sector of economy (Ukraine)

Modelación del ciclo del proceso de reproducción en el sector agrario de la economía (Ukrania)

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Received: 31/10/2018 • Approved: 12/02/2019 • Published 04/03/2019

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ABSTRACT:

The methodological basis was investigated and the use of periodogram analysis, Stokes' integrals, Buys-Ballot scheme and autocorrelation functions for the analysis of cyclicity in reproduction processes was suggested. By analyzing the dynamics of the share of capital investments in gross value added for the period of 1956-2017, long-term, medium-term, short-term cycles and micro-cycles of reproductive processes in the national economy and the agrarian sector of the economy were revealed.

Keywords: cyclicity, reproduction process, hidden periodicities, agrarian sector of economy.

RESUMEN:

Es investigada la base metodológica y el uso del análisis de periodogramas, Stokes integrales, esquemas de Buys-Ballot y las funciones autocorrelativas para el análisis del ciclo de los procesos de reproducción. Por medio del análisis de la dinámica de las inversiones capitales en el valor añadido bruto en 1956-2017. Son revelados a largo, mediano y corto plazo, los ciclos y los microciclos de los procesos de reproducción en la economía nacional y el sector agrario de la economía.

Palabras clave: ciclicidad, proceso de reproducción, periodicidad oculta, sector agrario de la economía

1. Introduction

The economic life of society is organically connected with the processes of reproduction of all elements of the economic system: economic relations, their subjects, productive forces and economic mechanism. In its turn, economic development is characterized by a periodic change in the process phases, which is based on the cyclicity of the reproductive process.

Currently, the agrarian sector of the economy of Ukraine is a system-oriented segment of the national economy, where on average 13.4 % of gross value added was created in 2015-2017. The food security of the State, employment in the rural areas and in the processing

industry, tax and other revenues to the budget depend heavily on the stability of its development. However, fluctuations in economic and market conditions, changes in weather and climatic conditions, the transformation of institutional support for entrepreneurship, globalization challenges and the dynamics of global trends in economic development lead to instability. Cyclicity of reproductive processes in the agrarian sector of the national economy leads to the emergence of a significant number of risks.

One way or another, the fulfillment of the tasks set by the Cabinet of Ministers of Ukraine related to the provision of the sustainable development of the agrarian sector and rural areas based on an innovative basis, to the guaranteeing of the country's food security, preventing over-production crises requires a well-considered countercyclical policy. In its turn, the effective anti-cyclical measures should be based upon the scientifically grounded economic and mathematical models. The allow to estimate the hidden periodicity of economic processes, to predict them in the short-and long-term perspective.

2. Methodology

2.1. Framework

The theory of economic cycles, alongside with the theory of economic growth, belongs to the theory of economic dynamics, which explains the movement and the development of the national economy of a country. In economic theory, there are about two hundred concepts and theories of the causes of crises and their cyclicity. At the same time, up to this time, the nature of the cycle remains one of the most controversial and little-studied issue.

The industrial crises of the XIXth century, as well as the documented experience of previous agrarian crises, led to the organized research of these issues conducted by the classics of economic thought and, thus, to the formation of the theory of business cycles. Agrarian cycles were studied mainly from the point of view of fluctuations in prices for agricultural products, the dynamics of yields and the influence of solar activity on them. A detailed analysis of the genesis of cyclical economic development is given in the works (Aimar, Bismans & Diebolt, 2016; Pustovoit, 2016; Podlesnaya, 2017). One of the first computational schemes aimed to detect hidden periods in agroeconomic and agrobiological research is the method of linear selective transformations, developed by C. Buys-Ballot in 1847. According to him, if the trial period coincides with the actual period of the investigated process, then averaging on the segments of the initial time a series of length equal to the trial period does not change the value of the global period, while smoothing the components of other periodic perturbations (Buys-Ballot, 1847). On the other hand, the Buys-Ballot transformation is equivalent to the product of the output discrete process and the periodic sequence of impulses with the subsequent averaging of this value (Serebrennikov & Pervozvanskiy, 1965, p. 38).

It can be argued that the problem of detecting hidden periodicity in a discrete time- series belongs to the class of incorrectly set. In general, it reduces to the finding of such transformations of the source data so that it is possible to determine the parameters of the periodic components. In 1898, for the first time, A. Schuster suggested a comprehensive solution to this problem, when he substantiated the scheme of periodogram analysis (Schuster, 1898). In addition, the scientist introduced the notion of a phase diagram, which makes it possible to more accurately determine the frequencies that are approximated by the abscissa of extremums of amplitude diagrams. The use of phase diagrams in the periodogram analysis of economic processes simplifies the decision-making regarding the choice of this or that trial period as a global period of the seasonal component.

As a result of the research, O. Oliinyk also concluded that the most universal function for analyzing the cyclical vibrations of economic processes is the sinusoid, while accounting for the linear trend (Oliinyk, 2005, p. 67). Further studies of the cyclicity of the main reproductive processes in agriculture enabled O. Oliinyk to show that the linear trend can be replaced by any other trend equation, while the number of sinusoidal harmonics is advisable to limit to one.

According to literary sources, there are now more than 30 methods and approaches allowing to detect hidden periods (Serebrennikov & Pervozvanskiy, 1965; Kufenko, 2016). The main ones are linear transformations (Brooks' transformation, Buys-Ballot scheme, ordinates averaging for the trial period), nonlinear selective transformations of poliharmonic processes (correlation methods, splines), periodogram analysis, Stokes' integration method, singular spectral analysis, etc. In particular, we conducted a comparative analysis of the methods of periodogram analysis, Stokes' integrals, autocorrelation functions, and Buys-Ballot analysis in terms of their relevance for detecting hidden periods in discrete time-series and assessing the cyclical nature of economic processes (Hutorov & Pogorelov, 2006; Hutorov, 2009). In addition, it was proved that in agroeconomic studies it is expedient to use several methods with further comparison and a critical analysis of the obtained results. This allows to define the durations of the global period and periods of higher order harmonics more accurately and with higher quality.

2.2. Economic aspect of simulation of cyclicity in the reproduction process in the agrarian sector of the economy

Some of the first characteristics of the reproduction cyclicity were provided in works of D. Ricardo. In particular, he noted that at different stages of society development, the capital accumulation has different speeds and depends on the productive forces of labor (Ricardo, 2001, p. 61). Not allowing the possibility of a crisis of overproduction, D. Ricardo did not fully consider the features of reproductive processes in the economy. At the same time, he generally justified the tendency of profit margins to decrease, thereby laying the foundations for K. Marx theory of industrial cycles.

Afterwards, in the early 1860's, C. Juglar showed the existence of medium-term economic cycles, lasting for 8-10 years. Their occurrence was determined by fluctuations in the volume of investment in the active part of fixed capital, the emergence of new types of technology (Juglar, 1862). At the same time, the increase in prices for capital goods leads to economic growth, and the decrease - to decline. In modern interpretation the very business cycle of Juglar has four phases (Podlesnaya, 2017, p. 41): recovery, expansion, recession and depression.

Thoroughly examining the capital movement, the mechanisms of its reproduction and the genesis of industrial crises in England, K. Marx concluded that the material basis of the cyclicality of the capitalist crises is the periodic renewal of fixed capital (Marx, 1969, p. 464-465). The duration of the business cycle suggested by K. Marx is 10-11 years, which is in general quite comparable with the results of C. Juglar.

In the context of studying the cyclicity of the dynamics of reproductive processes, the scientific contribution of M. Tugan-Baranovsky looks reasonable. Having analyzed different approaches to the essence of economic cycles, he concluded that the obstacle to sustainable economic and sectoral development is not so much exogenous factors, as the endogenous bifurcations of the economic system of different levels of the hierarchy, which also give rise to non-stationary processes. M. Tugan-Baranovsky, thus, proved the existence of cyclic laws in reproduction of the country's fixed capital (Tugan-Baranovsky, 1914, p. 285-305). In addition, the scientist has shown that the production cycle is explained not only by the periodic renewal of fixed capital, but also by the dynamics of market prices for the means of production and consumer goods. He explained the periodic capitalist crises by the inertia of investment activity, thereby outstripping J. Schumpeter (Schumpeter, 1939).

While considering various factors of reproduction, J. Kitchin in 1923 substantiated the microcycles of economic dynamics lasting for 40 months (Kitchin, 1923). They are related to the circulation of material goods, however, unlike the cycles of Juglar and Marx, they are based on the reproduction of current assets, the movement of inventories, the investment lag, caused by managerial decisions, information asymmetry, and so on. In addition, during the Kitchin cycle, there is no renewal of fixed capital.

In contrast to the cycles of reproduction of the active part of fixed capital and working capital, S. Kuznets modeled the periodicity of medium- and long-term economic dynamics,

lasting for 15-20 years, through the reproduction of the passive part of capital goods (civil and industrial construction) (Kuznets, 1930). In his opinion, a function that describes the cyclic reproduction process should: have a certain boundary; reflect the decrease in the percentage rate of total output increase; to reflect the change in the absolute value of output due to influence of stimulating and de-stimulating factors; to show a decrease in the pace of growth while reaching the limit (Kuznets, 1930, p. 64). From the set of functions describing growth, S. Kuznets chose the Pearl-Reed (1) and Gompertz (2) curves:

$$O(t) = \frac{L}{1 + e^{a-bt}}, \quad b > 0, \tag{1}$$

$$O(t) = ce^{a(1-R^t)}, \ 0 < R < 1,$$
 (2)

where L – the marginal volume of the industry production; t is a time factor; a, b, c – the estimated parameters of the function O(t); R – ratio of actual output to the marginal one: $R = O_t/L$, t = 1, ..., n; n – number of time periods under investigation.

According to S. Kuznets, equation (1) characterizes the logistic curve, which tends to decrease due to increased competition in capital markets. In addition, the Pearl-Reed and Gompertz curves are suitable for modeling both the growing and decreasing trends in the volumes of industry output.

Basing on the results of K. Marx and M. Tugan-Baranovsky, M. Kondratiev proved the existence of long cycles of economic dynamics, lasting for 40-50 years. To do this, the scientist analyzed over 100 years of the time-series of the main indicators of the national economy development: national capital (K), the number of able-bodied population (A), the production of means of production and means of consumption, total national production, gross domestic product (E), nominal wages (I), interest on capital (i), land rent, size of capital accumulation (S) (Kondratiev, 2002, p. 503). Taking into account that national capital, the number of able-bodied population and the level of technical and technological development of the economy are stock values, while other characteristics of the national economy are the flow ones, M. Kondratiev determined the general trend equations of their changes. Thus, the general trend of cumulative indicators is determined by the differential equation (3).

$$\partial y/\partial t = ky(L-y),$$
 (3)

where $y: M \subset {}^{\sim} \to {}^{\sim}$, $\forall t \in M, y \Rightarrow y(t)$; k – proportionality coefficient; L – parameter of change of a certain cumulative indicator.

The solution of equation (3) is function of a Pearl-Reed type:

$$y(t) = \frac{L}{1 + Ce^{-akt}},\tag{4}$$

where L, C, and α are the desired parameters of the function y(t).

Then M. Kondratiev built a differential equation in partial derivatives, which shows the relationship between gross domestic product, national capital and the number of able-bodied population (Kondratiev, 2002, p. 505):

$$E = \frac{\partial E}{\partial K} K + \frac{\partial E}{\partial A} A \Longrightarrow E = m\sqrt{AK}, \tag{5}$$

where m is the average level of technical and technological development of the national economy (technological coefficient, which is got from the assumption that GDP elasticity by national capital is equal to GDP elasticity by population).

At the same time, formula (5) implies that $\partial E/\partial A=l$ and $\partial E/\partial K=i$. It reflects the dynamics of the average

interest on capital and of the average nominal wage in the national economy. In addition, provided that $m=\sqrt{2il}$, the

value of the aggregate national capital is equal to the value of cumulative labor: iK = lA, and the average level of labor productivity is determined by the ratio of average nominal wage to average interest on capital: in the classical economic theory this determined the directions of anti-cyclical regulation.

M. Kondratiev determined the rate of reproduction of capital in the national economy by to the dynamics of his accumulation (Kondratiev, 2002, p. 505):

$$S = \partial K/\partial t$$
, (6)

where the function of the national capital K(t) is determined by the equation (4).

Considering the identical nature of the reproducing cycles of Kitchin, Juglar, Kuznets, and Kondratiev, scientists see organic ties between them: 2 Kitchin's cycles form a 1 cycle of Juglar; 2 cycles of Juglar – 1 cycle of the Kuznets; 2 cycles of the Kuznets – 1 long cycle of Kondratiev (Filin, 2015, p. 21).

The dynamics of capital reproduction (6) is objectively determined by the change in the volume of capital investments (1) that has been mastered over the period t. In general, this idea was thoroughly considered by J. Keynes and his followers. Thus, J. Keynes proved that the dynamics of demand for investment (investment impulses), which nonlinearly affects the change in

aggregate demand in national economy, is the reason of economic cycles (Keynes, 2012). In turn, the motive for investing is to generate income on invested capital. The combination of these preconditions enabled the Keynesians to justify the concept of an investment multiplier, which shows a non-linear relationship between the volume of capital investment and national income, and determines the vector of its changes. On the other hand, the presence of induced investment is explained by the size of national income, which causes a change in demand for them (the effect of an accelerator). Thus, the mechanism of the accelerator multiplier characterizes the nature of the economic cycle caused by the social reproduction process, and also explains the dynamics of the economic system in time.

Based on the Harrod-Domar economic growth model and the works of J. Keynes, the Polish economist M. Kalecki modeled the cyclical nonlinear dynamics of the reproductive investment process (Kalecki, 1935, p. 332):

$$\frac{\partial I}{\partial t} = \frac{m}{\theta} I(t) - \left(\frac{m}{\theta} + n\right) I(t - \theta) + nU,\tag{7}$$

where I(t) — nonlinear function of the capital investments change in time; t is time factor; m — technological coefficient, indicating the level of innovation development of the economy, m > 0; ϑ — investment lag; n — the level of depreciation of fixed capital, $n \ge 0$; U — constant amount of demand for the reproduction of fixed capital.

In general, the solution for the differential equation (7) will be a function of the form $I(t) = C_1 e^{-C_2 nt} + U$, where C_1 , C_2 are constants. Applying the Euler formula to it, we will get a system of complex numbers in a trigonometric form. Then, it is also correct to expand it into the Fourier series, whose parameters are determined by the Euler-Fourier formulas. Similar transformations are valid for equations (1), (2) and (4) as well.

Making similar transformations, M. Kalecki showed that the period of the cycle of the reproduction process, described by the differential equation in differences (7), depends on the investment lag, on the innovative development of the national economy and on the average depreciation of fixed capital (Kalecki, 1935, p. 335, 337):

$$T = \frac{2\pi}{\arccos\left(\frac{m}{m+n\theta}\right)}\theta. \tag{8}$$

In 1935, scientists R. Frisch and H. Holme proved that equation (7) has a unique solution - a sinusoidal function with a period $T=2\pi\theta/u$, where u is the real part of the complex number obtained after applying Euler's formula (Frisch & Holme, 1935). This solution completely coincides with the conclusions of M. Kalecki.

2.3. Methods of estimating hidden periods in the dynamics of the reproduction process in the agrarian sector of economy

Assume that the function f(t) contains four terms: the periodic component p(t), the non-stationary trend u(t), different from p(t), and the white noise ε ($\varepsilon \sim N(0, \sigma_{\varepsilon}^2)$):

$$f(t) = p(t) + u(t) + \varepsilon. \tag{9}$$

Note that the periodic component ho(t) is characterized by the global period au and frequency $\omega=2\pi/T$; it is characterized

by the equality p(t+T) = p(t), $\forall t \in M, T \neq 0$. The non-stationary trend u(t) can also contain hidden periodic

components of low power, the characteristics of which are not necessarily different from the parameters of periodicity p(t). It is well known that any periodic function meeting the Dirichlet conditions can be expanded into a Fourier series on the interval [-T, T], with a period equal to the length of the interval. However, in economics the problem of decomposition of the final realization of the process on the orthogonal basis of trigonometric functions is methodologically incorrect, because the sampling interval is not constant. On the other hand, spectral analysis requires the presence of a strictly periodic function, defined at an infinite interval of time. Since discrete economic reproduction processes are not infinite, their exact time of beginning and ending is unknown, then the periodic extension of quasiperiodic processes to an infinite domain of values with a certain period of the fundamental harmonic T leads to a shift in estimates, to the flow of the power of the periodic process to the harmonics of higher orders at the points of forced ruptures, distortion in the economic meaning of the obtained results,

In this relation, there is a need for other methods of studying hidden periods in the dynamics of the reproductive process in the agrarian sector of the economy, which are different from the spectral Fourier analysis. Among such methods, we have substantiated the use of periodogram analysis. Stokes integrals the Buys-Ballot method and autocorrelation functions.

(Hutorov & Pogorelov, 2006; Hutorov, 2009).

The method of periodogram analysis is aimed at approximating of a discrete time-series by a trigonometric function, basing on the criterion of the minimum of the mean square deviation. Then, based on the trigonometric form of the solution for the differential equation in differences (7), the approximating function V(T) will have the following form:

$$V(T) = C_0 + \sum_{i=1}^{K} A_i \cos \frac{2\pi i}{T} t + B_i \sin \frac{2\pi i}{T} t,$$
 (10)

where C_0 , A_i and B_i are the estimated parameters; K – the specified number of harmonics.

The condition of the best approximation will be the minimum of the functional

is more resistant to the Slutzky-Yule effect than the classical periodogram analysis.

$$R\left(T_{p}\right) = \sum_{i=1}^{N} \left(t_{i} - C_{0} - \sum_{j=1}^{K} \left(A_{j} \cos \frac{2\pi j}{T_{p}} t_{i} + B_{j} \sin \frac{2\pi j}{T_{p}} t_{i}\right)\right)^{2}, \text{ where } T_{p} \text{ is the trial period. The value of } T_{p} = \sum_{i=1}^{N} \left(T_{p} - \sum_{j=1}^{K} \left(A_{j} \cos \frac{2\pi j}{T_{p}} t_{i} + B_{j} \sin \frac{2\pi j}{T_{p}} t_{i}\right)\right)^{2}$$

the global period T is determined via iteration on the interval of the trial periods T_{ρ} = 1,...,L, $L \le N$. For this, it is necessary to solve (2L + 1) equations with the same number of unknowns given $L \le T_{\rho} / \Delta t_{\rm max}$, where $\Delta t_{\rm max}$ – the maximum sampling interval of the output timeline (Antomonov, 1983, p. 116). The criterion for the quality of approximation will be the

value
$$\gamma = 1 - \frac{R(T)}{GSS}$$
, where GSS – the total sum of squared deviations of the output time-series from the average one

(Antomonov, 1983, p. 112). In addition, γ characterizes the power of the periodic component p(t) with the global period T, and the function $\gamma(T)$ is the periodogram of the reproduction process.

The method of Stokes integrals is a simplified kind of periodogram analysis. The approximating function $V(T_p)$ has the following form:

$$V(T_p) = \left[\left(\frac{1}{N-1} \sum_{i=1}^{N} \sum_{j=1}^{K} t_i \sin \frac{2\pi j}{T_p} \right)^2 + \left(\frac{1}{N-1} \sum_{i=1}^{N} \sum_{j=1}^{K} t_i \cos \frac{2\pi j}{T_p} \right)^2 \right]^{\frac{1}{2}}. (11)$$

The values of the sinusoidal and cosine components in (12) coincide with the analogous values in (10), with the accuracy to the normalization factor 1/(N-1), where (N-1) shows the velocity of the time-series. The periodogram $V(T_p)$ is only an indicator of the presence of periodic components, which makes it impossible to assess their reliability. Therefore, the choice of the global period T in the interval of the trial periods $T_p = 1,...,N-1$ is also determined by the iterative method according to the same criteria as for the periodogram analysis. Considering the form $V(T_p)$ of the functional, the method of Stokes integrals

The Buys-Ballot method is a linear selective transformation of the output time-series for evaluating the value of the global period both in p(t) and in u(t). According to this method, the time-series investigated is divided into sections multiples of the trial period: $N = rT_p$, where r is an integer. Then, it is averaged over these intervals. Approximation function as such is absent. The general view of the transformations according to the scheme of Buys-Ballot is given as follows:

$$Q_{k}(t) = \frac{1}{2k} \sum_{i=-k}^{k-1} f\left(t + \left(i + \frac{1}{2}\right)T_{p}\right), k = \frac{1}{2}r,$$

$$Q_{k}(t) = \frac{1}{2k+1} \sum_{i=-k}^{k} f\left(t + iT_{p}\right), k = \frac{r-1}{2},$$
(12)

where k is an integer.

We showed (Hutorov & Pogorelov, 2006, p. 86) that the criterion for evaluating the linear selective transformation of the Buys-Ballot is $D_k = \max Q_k - \min Q_k$, k=1,...,r. The greater the value of D_k in the definition area, the more the trial period T_p approaches the real period of the investigated economic process. In addition, $\left\{D_k \mid k \in 1...r\right\}$ can be considered as the "periodogram" by the Buys-Ballot, and the value of the global period is defined as $T = \max D_k$.

According to the Buys-Ballot method, the frequency characteristics is a periodic function of $H(\psi)$ with a period $\psi=2\pi/T$ (Serebrennikov & Pervozvanskiy, 1965, p. 36):

$$H(\psi) = \frac{\sin r \frac{1}{2}}{r \sin \frac{\psi T}{2}}.$$
 (13)

Obviously, the highest maxima $H(\psi)$ are located at points $\psi=2\pi l/T$, where l is an integer that generally corresponds to a plurality of values $\left\{D_k \mid k \in 1...r\right\}$. In addition, the significance of the maxima $H(\psi)$ increases with an increase in the value of r.

Subsequently, the value of the global period T got after applying the Buys-Ballot method is used to construct the function of the form (10) and to conduct a periodogram analysis on its basis.

Another method of detecting hidden periods in discrete time-series of economic processes is auto-correlation functions. Thus, in 1937 E. Slutzky proved that empirical correlation functions are pseudo-periodic, tend to sinusoidal form with increasing number of the time-series members (Slutzky, 1937). In addition, according to the Wiener-Khinchin theorem, the auto-correlation signal function is a Fourier transformation of its spectral density of power, which also indicates its periodic nature. In general, the autocorrelation function $K(f, \tau)$ shows the dependence between the output function f(t) and its copy $f(t+\tau)$, shifted by a certain value τ . Proceeding from the essence of the correlation, it is obvious that the autocorrelation function is symmetric with respect to the moment $\tau = 0$: $K(f, \tau) = K(f, -\tau)$, and therefore can only be determined for positive values of τ , which is quite sufficient for the analysis of reproductive processes. The function $K(f, \tau)$ reaches its maximum at zero: and the minimum for

$$\tau \to \infty$$
: min $K(f,\tau) = K(f,\infty) = 0$.

The determination of hidden periodicities using autocorrelation functions is since $K(f, \tau)$ represents a new time-series whose values are more correlated than in the initial series f(t). Then the periodic components in f(t) become more pronounced in $K(f, \tau)$ as periodic with the parameter τ . At the same time, the autocorrelation function does not contain phase information, but allows one to estimate only the period and frequency of the process without an approximation of the discrete time-series. We set the discrete autocorrelation function $K(f, \tau)$:

$$K(f,\tau) = \begin{cases} \frac{\Theta}{\sqrt{Z}}, Z > 0\\ +1, Z \to 0, \Theta \ge 0,\\ -1, Z \to 0, \Theta < 0 \end{cases}$$
(14)

$$\text{where } \Theta = \sum_{i=1}^{N-\tau} t_i t_{i+\tau} - \frac{\sum_{i=1}^{N-\tau} t_i \cdot \sum_{i=1}^{N-\tau} t_{i+\tau}}{N-\tau} \; ; \; Z = \left(\sum_{i=1}^{N-\tau} t_i^2 - \frac{\left(\sum_{i=1}^{N-\tau} t_i\right)^2}{N-\tau}\right) \cdot \left(\sum_{i=1}^{N-\tau} t_{i+\tau}^2 - \frac{\left(\sum_{i=1}^{N-\tau} t_{i+\tau}\right)^2}{N-\tau}\right) \cdot \left(\sum_{i=1}^{N-\tau} t_i^2 - \frac{\left(\sum_{i=1}^{N-\tau} t_i\right)^2}{N-\tau}\right) \cdot \left(\sum_{i=1}^{N-\tau} t_i\right) \cdot \left($$

Rationing of the autocorrelation function using the dispersion ($K(f,\tau)/\sigma_t^2$) suggested by G. Jenkins and D. Watts (Jenkins & Watts, 1968, p. 146), we consider as inappropriate in our case. Then, on the displacements interval $\tau \in [1, N-2]$, we will sequentially determine the values of the autocorrelation functions (14):

 $\left\{K\left(f,\tau\right)|\,\tau\in 1...N-2\right\}. \text{ Using the iterative method, we determine the maximum }\mathit{K}(\mathit{f},\tau) \text{ over the displacement interval }\tau\in [1,N-2]. \text{ Then the value of the global period }\mathit{T} \text{ for the time-series studied will be }\tau \text{ at the point of maximum }\mathit{K}(\mathit{f},\tau)\text{: }T=\underbrace{\tau}_{\max K(f,\tau),\tau\in[1,N-2]}.$

2.4. Research materials

Based on the above stated, we took the share of capital investment in gross value added in percentage terms as an indicator for studying the cyclicity of the reproduction process.

We fully share the opinion of M. Allais, the longer the study period (the "memory of the economic process"), the stronger and more stable the trend is, and the longer the global period is (the period of the main cycle) (Allais, 1998, p. 124). Our research covers actual data on Ukraine for 1956-2017, its source is the statistical yearbooks of the Central Statistical Office under the Council of Ministers of the Ukrainian SSR (1956-1986), the State Committee of the Ukrainian SSR on Statistics (1987-1990), the Ministry of Statistics of the Ukrainian SSR (1991), the Ministry of Statistics of Ukraine (1992-1996), the State Statistics Committee of Ukraine (1997-2009), the State Statistics Service of Ukraine (2010-2017). The data for 1918-1955 were not analyzed due to the peculiarities of the state creation of that time, the crises of the war and postwar years. This could've lead to a shift in the general trend and seasonal component. Information for 2014-2017 is given without taking into account the temporarily occupied territory and the anti-terrorist operation zone.

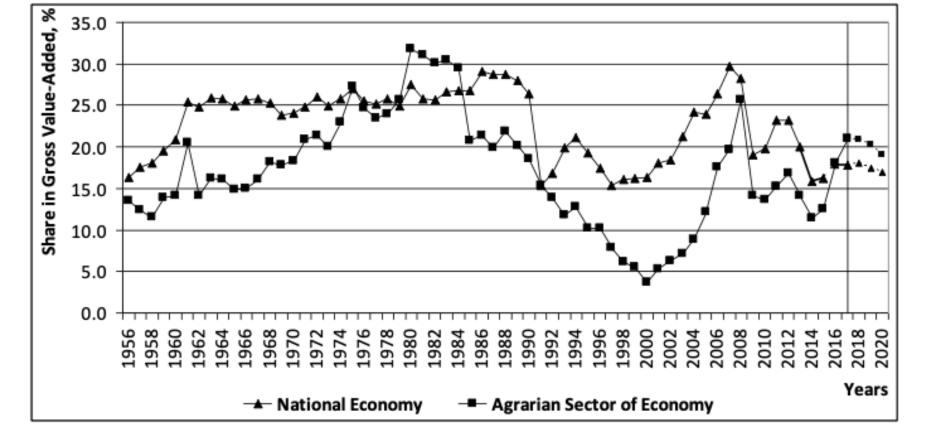
Gross value added for 1956-1990 was approximated by national income, taking into account the methodological comparability of these indicators. The data on the agrarian sector of the economy correspond to the "Agriculture branch" according to the All-Union Classifier of the branches of the national economy (1956-1991), the General Classifier "ndustries of the National Economy of Ukraine" (1992-1996), section "A" of the Classifier of Types of Economic Activities, 1996, and "Agricultural and Hunting, Forestry" (1997-2011), Section "A" of the Classification of Types of Economic Activities, 2010, "Agriculture, Forestry and Fisheries" (2012-2017). Since aggregations are generally the same, their analysis in dynamics is methodologically correct. No additional smoothing, filtering and interpolation of the time-series were conducted.

3. Results

The fixed capital accumulation in Ukraine in 1956-2017 was uneven, especially during the reforms of the mid-1990's. In particular, in 1995-2004, even a simple reproduction in most sectors of the economy was not ensured. The stabilization of the national economy after market transformations intensified the reproductive processes, which affected the annual growth of net accumulation of fixed capital (especially industrial) in 2005-2009. However, the recession in 2012-2013 and the crisis began in 2014 led to a decline in the level of investment attractiveness, led to the outflow of capital from Ukraine and to a new stage of decapitalization and de-industrialization (Hutorov, 2017, p. 188).

The basis of fixed capital accumulation and its reproduction is capital investment. The priority of the choice of investment objects usually depends on the level of return on investment, which, in normal business conditions, corresponds to the dynamics of the norm of value added. That is, under the conditions of neoindustrial development, the objects of higher technological redistribution are more investment attractive, because they create more value. However, the export orientation of the agrarian sector on raw materials, its disintegration with the processing industry deformed the institute of investments, causing an excess of capital investment rather into the production of raw materials than into the food industry. At the same time, these capital investments are unproductive, their economic value does not correspond to the cost, as indicated by their volumes exceeding the gross accumulation of capital assets of the agrarian sector in 2013-2017. In other words, a significant part of capital investment does not lead to an increase in the capitalization of the industry, they are not mastered properly, the fundamental principles of reproductive processes are undermined. Given that GDP consists of household consumption expenditures, government consumption and investments, gross private investments and net exports, the dynamics of the share of capital investment in gross value added also seems to be critical.

Output data are discrete time-series, containing 62 members, the graph of which is shown in Fig. 1



Notes. The data for the 1956-2017 are actual, for 2018-2020 – forecasted (in the figure marked with a dotted line after the vertical one).

Source: author' computations.

The discrete time-series of the share of capital investments in gross value added in Ukraine are characterized by a sufficiently large variation to put forward the hypothesis that they have periodic components. They are unstable relative to the average value (Tab. 1).

Table 1Descriptive statistics of the time-series of the share of capital investments in gross value added in Ukraine in 1956-2017

| Indicator | National economy | Agrarian sector of economy |
|-------------------------|------------------|----------------------------|
| Simple arithmetic mean, | 22.756 | 17.020 |
| Standard error, SE | 0.544 | 0.862 |
| Median, <i>ME</i> | 24.519 | 16.167 |
| Standard deviation, | 4.280 | 6.786 |
| Sample dispersion, | 18.317 | 46.047 |
| Asymmetry, β1 | -0.351 | 0.255 |
| Excess, β2 | -1.283 | -0.328 |

Source: author' computations

The absence of missed values allows the use general methods of statistical analysis without the need for additional interpolation. For data processing, we applied A. Hutorov's computer program "Expert system of "Periodograms", which contains advanced algorithms for the search of hidden periodicities. During the calculations, the following basic parameters were taken into account: the maximum number of trial periods – 1000; the minimum value of the global period is 2; the maximum value of the global period is equal to the number of members in a row.

As a result of calculations, the hypothesis on the existence of periodic components of different durations in time-series was confirmed (Tab. 2).

Table 2Characteristic of the periodic components in the time-series of the share of capital investments in gross value added in Ukraine in 1956-2017

| Method of analysis | The object of analysis | The number of periodic components | The values of periods (by the descending value of the component power) | Global Period, 7 |
|-----------------------------------|------------------------|-----------------------------------|--|---------------------|
| Periodo-gram Analysis | N. E. | 14 | {20.1; 56.4; 11.2; 14.4; 8.4; 6.4; 5.8; 4.5; 3.6; 7.3; 2.9; 4.0; 3.2; 2.4} | 20.1 |
| | A.S. | 15 | {47.0; 13.9; 11.5; 9.1; 4.6; 6.9; 5.2; 4.0; 5.7; 2.2; 7.8; 3.7; 3.1; 3.3; 2.7} | 47.0 |
| Stokes Integrals | N. E. | 8 | {42.0; 24.0; 62.0; 17.0; 12.0; 8.0; 4.0} | 42.0 |
| | A.S. | 8 | {43.0; 26.0; 62.0; 17.0; 14.0; 10.0; 6.0; 3.0} | 43.0 |
| Buys-Ballot | N. E. | 8 | {20.0; 18.0; 14.0; 12.0; 9.0; 6.0; 3.0; 28.0} | 20.0 |
| | A.S. | 9 | {26.0; 23.0; 28.0; 20.0; 18.0; 14.0; 12.0; 9.0; 4.0} | 26.0 |
| Auto- correlation Functions | N. E. | 6 | {41.0; 35.0; 22.0; 18.0; 14.0; 53.0} | 41.0 |
| | A.S. | 6 | {56.0; 53.0; 47.0; 45.0; 41.0; 37.0} | 56.0 |

Notes. N. E. – national economy; A.S. – is the agrarian sector of the economy *Source:* authors' computations.

Data of table 2 show that the reproduction process in the Ukraine has four distinct cycles: global (41 \pm 1 year), medium term (20 \pm 1 year), short-term (9 \pm 1 year) and micro-cycle (3 \pm 1 year). Mathematical models describing the periodic process are significant for Fisher's F-criterion at the confidence level of 0.99. Similar calculations for the dynamics of the share of capital investments in the gross value added of the agrarian sector of the economy show the existence of a global period lasting for 52 \pm 5 years, a medium-term cycle of 26 \pm 1 years, a short-term cycle of 13 \pm 1 years and a cycle of 6 \pm 1 years.

The decomposition of the time-series to the harmonics leads to a decrease in the statistical significance of the members of the higher orders harmonics as their number increases. In addition, this also causes the appearance of "parasitic" maxima on periodograms due to the displacement of the spectral density of power.

The period values got generally correspond to the normative terms of the use of fixed assets, while indicating the optimal time periods for reproduction cycles. In the agrarian sector of the economy, the increase in the average cycle duration is explained by the peculiarities of the reproduction of long-term biological assets, the influence of solar cycles on agricultural production, and so on.

To predict reproductive processes under the research, we used neural networks available in the StatSoft Statistica v. 10. The base neural network has been chosen a multilayered perceptron with architecture (6-2-1), with the Quasi-Newtonian algorithm for training of Broyden-Fletcher-Goldfarb-Shanno (BFGS), by the sinusoidal activation function of the neurons source and the identical activation function of the hidden neurons, as well as the

projection window for 8 years. As a result of the training of 20 thousand neural networks during 20 thousand eras for each, two were selected with errors of 14.23 % and 15.89 % respectively. Besides, the control errors are 9.33 % and 9.87 %, which indicates a high level of reliability of 8-year forecasts. Forecasted in this way data for 2018-2020 shows that the short-term cycle with the ascending wave has been completed in 2017. Then, the period of investment and reproductive processes decline is expected both in the national economy, as a whole, and in the agrarian sector of the economy.

4. Discussion of results

The resulting simulation cycles of the reproductive process in the Ukrainian economy and its agrarian sector generally correspond to the results of other scientists.

Thus, V. Podlesnaya gives data on the medium-term economic cycles lasting about 20 years (Podlesnaya, 2017, p. 324), as well as on long cycles of the reproductive process lasting 45-54 years (Podlesnaya, 2017, p. 323). The latter are caused by changes in technological patterns and the introduction of epoch-making innovations. The development of the national economy of Ukraine is characterized by the existence of micro-cycles, lasting 3-5 years (Podlesnaya, 2017, p. 419). According to O. Pustovoit, during 1997-2015 medium-term institutional cycle had place in Ukraine, it consists of two short-term cycles of 9 years (Pustovoit, 2016, p. 355).

The study of the crop yields dynamics in 1795-2007 and that of gross agricultural output in 1909-2006 allowed V. Rastyannikov and I. Deryugina to detect long 50-60 years cycles (Rastyannikov & Deryugina, 2009, p. 111).

According to I. Sokolov, I. Tarapatov and O. Borysenko, the cyclicality of grain production in Ukraine has the 14-years period (Sokolov, Tarapatov & Borysenko, 1996, p. 43). The results of the analysis of the dynamics of grain yields, sunflower seeds and sugar beets conducted by O. Oliinyk showed that they demonstrate short-term cycles of 15-18 years and microcycles of 2-6 years. At the same time, according to the scientist the cyclicity of solar activity is described by 10.5-11.0 years period (Oliinyk, 2005, p. 285).

According to I. Zagaytov, the duration of short-term reproduction cycles in the agrarian sector of the economy (Tc) is a function f(Ka) of the duration of the renewal of the active part of fixed capital (Ka): where ΔD is the growth rate of real solvent demand in the intercrisis period; ΔPr – increase in the productivity of the renewed capital (Zagaytov, 2011, p. 81-82). Thus, the scientist defined the average duration of short-term cycles at the level of 13 years.

5. Conclusions

The cyclicity of the reproduction process is an integral part of its dynamics, characterizing the evolutionary course of the development of productive forces, economic relations and economic mechanism within a country under the influence of global social transformations. These cyclical fluctuations are usually caused by natural-economic, socio-economic and institutional factors, and their material basis is the processes of social reproduction of capital in its broad sense. Therefore, all the "classic" economic cycles are organically linked with the investments and innovations, and a periodic innovation renewal is a necessary for the development of society and economy, in the long run it marks a change in technological modes. In our opinion, the micro-cycles of the reproduction process (as well as the Kitchin cycles) are due to the investment lag, while short-, medium- and long-term cycles (as well as the corresponding cycles of Juglar, Marx, Kuznets, and Kondratiev) – to the dynamics of capital investments and the renewal of the main productive capital.

The results of critical analysis of the methodological basis of modeling the reproduction process cyclicity in the agrarian sector of the economy substantiated the use of periodogram analysis, Stokes integrals, and Buys-Ballot and autocorrelation functions for the estimation of hidden periodicity in discrete time-series. On the basis of these methods, the analysis of the dynamics of the share of capital investments in gross value added in Ukraine for 1956-2017 determined the duration of long-, medium-, short-term cycles and micro-cycles in both

the general reproduction process and in the agrarian sector of the economy. The obtained durations of cycles are statistically significant, closely correlate with the results of studies of other scientists, which confirm their reliability.

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Revista ESPACIOS. ISSN 0798 1015 Vol. 40 (Nº 07) Year 2019

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